

# Bayesian, Fiducial, and Frequentist (BFF): Best Friends Forever?

BFF 1/21

Xiao-Li Meng

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Replication!

Basu Ex

Summary

Xiao-Li Meng

Department of Statistics, Harvard University

- Liu & Meng (2106) **There Is Individualized Treatment. Why Not Individualized Inference?** *Annual Review of Statistics and Its Application*, 3: 79-111
- Liu & Meng (2014). **A Fruitful Resolution To Simpson's Paradox via Multi-Resolution Inference.** *The American Statistician*, 68: 17-29.
- Meng (2014). **A Trio of Inference Problems That Could Win You a Nobel Prize in Statistics (if you help fund it).** *In the Past, Present, and Future of Statistical Science (Eds: X. Lin, et. al.)*, 535-560.



# What is *inference*? Katie's answer ...



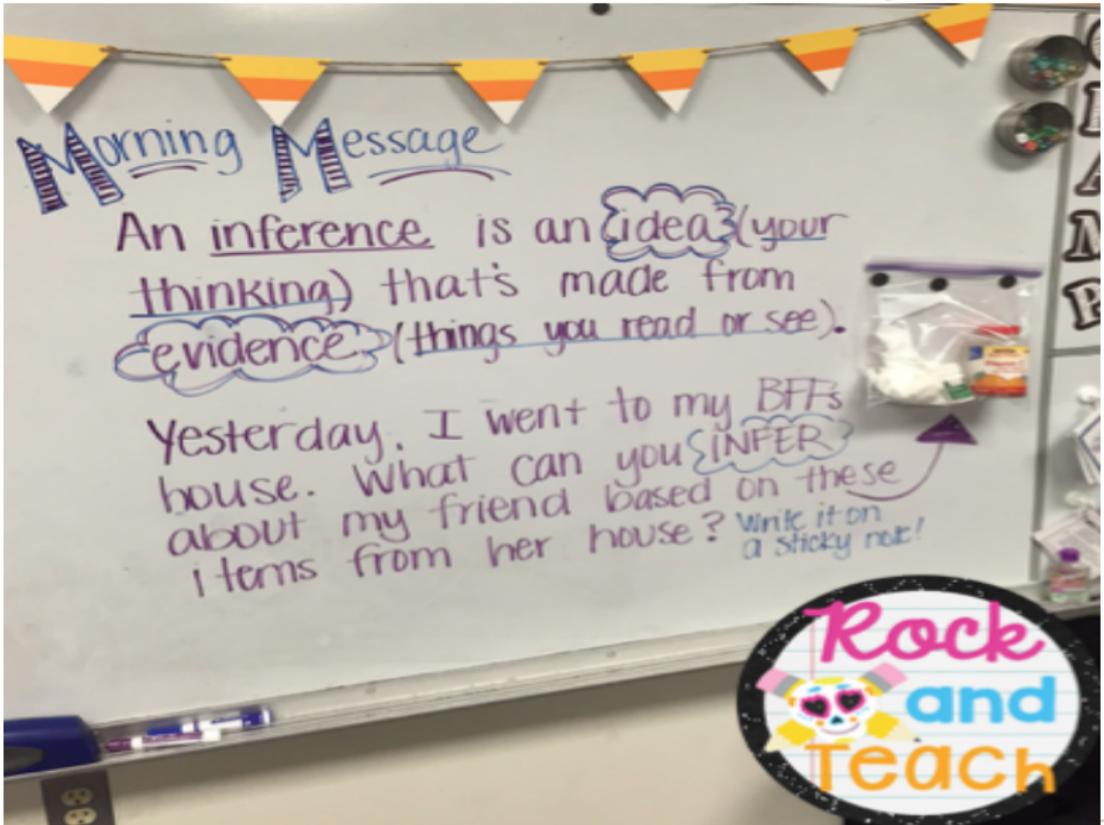
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## Morning Message

An inference is an idea (your thinking) that's made from evidence (things you read or see).

Yesterday, I went to my BFF's house. What can you INFER about my friend based on these items from her house? Write it on a sticky note!



# But what is *Statistical/Probabilistic* Inference?

- An ultimate intellectual game: **“to guess wisely and to guess meaningfully the errors in our guesses.”**  
(*XL-Files*, Oct 2015)

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(*XL-Files*, Oct 2015)
- Impossible to access exact errors, but a **full spectrum** of possibilities for accessing **probabilistic errors**.

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Pure Frequentist (Fully unconditional)

Most Robust but Least Relevant

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Pure Frequentist (Fully unconditional)

Most Robust but Least Relevant

Pure Bayesian (Fully conditional)

Most Relevant but Least Robust

But life is about *compromise*:

Conditional frequentist, Objective Bayesian, Fiducial ...



# It all depends on which *Replications* you want ...

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Summary

# It all depends on which *Replications* you want ...

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## Statistical Model via Stochastic Representation

$$\underbrace{D}_{\text{Data}} = G(\underbrace{\theta}_{\text{Signal}}, \underbrace{U}_{\text{Noise}}) \quad (S)$$

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Ex:  $D = \{X_1, \dots, X_n\}$ , where

$$X_i = \theta + U_i, \quad U_i \stackrel{\text{iid}}{\sim} N(0, 1),$$

and  $U = \{U_i, i = 1, \dots, n\}$  represents “God’s Uncertainty”

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- Frequentist: Fix parameter  $\theta$ , vary  $D$

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- Frequentist: Fix parameter  $\theta$ , vary  $D$
- Bayesian: Fix data  $D$ , vary  $\theta$
- Fiducial: Fix neither, but vary  $U$ , subject to the constraint (S) (or implied constraints with  $A(U)$  fixed)



# The differences are in the replications ...

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Summary

$\theta$   
|  
 $D$

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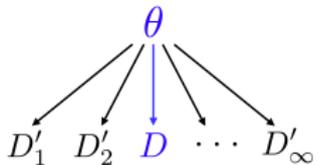
Summary

$$\theta$$

$$\downarrow$$

$$D$$

Frequentist Inference



$$p(D'|\theta)$$

# The differences are in the replications ...

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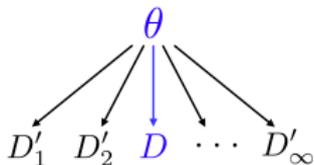
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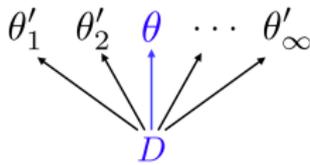


Frequentist Inference



$$p(D'|\theta)$$

Bayesian Inference



$$p(\theta'|D)$$

$$\propto p(D|\theta')\pi_0(\theta')$$

# The differences are in the replications ...

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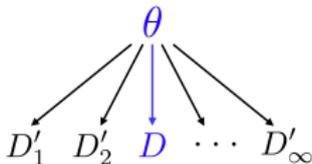
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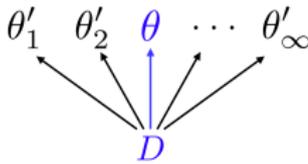


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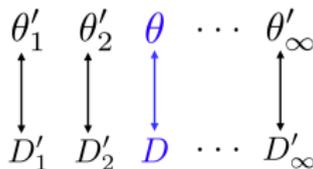
Bayesian Inference



$$p(\theta'|D)$$

$$\propto p(D|\theta')\pi_0(\theta')$$

Fiducial Inference



$$p(D', \theta'|A(U))$$

$$= p(D'|\theta', A(U))\pi(\theta')$$



# Illustrate BFF for $X \sim N(\theta, 1)$

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Summary

Frequentist



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Summary

Frequentist

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# Illustrate BFF for $X \sim N(\theta, 1)$

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Replication!

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Summary

Frequentist

Bayesian

Fiducial

# Illustrate BFF for $X \sim N(\theta, 1)$

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Replication

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Summary

Frequentist

Bayesian

Fiducial

Sampling Dist.

$$X|\theta \sim N(\theta, 1)$$

# Illustrate BFF for $X \sim N(\theta, 1)$

Frequentist

Sampling Dist.  
 $X|\theta \sim N(\theta, 1)$

Bayesian

+ Prior Dist.  
 $\pi_0(\theta) \propto 1$

Fiducial

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 $X - \theta = U \sim N(0, 1)$

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Confidence Interval

$$(X - z_p, X + z_p)$$

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Posterior Dist.

$$\theta|X \sim N(X, 1)$$

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$$X - \theta = U \sim N(0, 1)$$

Fiducial Dist.

$$\theta = X + U \sim N(X, 1)$$

Confidence Interval

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Basu Ex

Summar

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Sampling Dist.

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Sampling Dist.

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Confidence Dist.

$$N(X, 1)$$

Confidence Interval

$$(X - z_p, X + z_p)$$

Generate for  $i = 1, \dots$

$X_i|\theta \sim N(\theta, 1)$ , then  
 $(X_i - 1.96, X_i + 1.96)$   
 covers  $\theta$  95% of times

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Generate for  $i = 1, \dots$

$\theta_i \sim \mathbf{any} \pi(\theta)$ , &  
 $X_i|\theta_i \sim N(\theta_i, 1)$ , then  
 $(X_i - 1.96, X_i + 1.96)$   
covers  $\theta_i$  95% of times

# Finding the Right “Control Population”: Treating Data as Your Patient



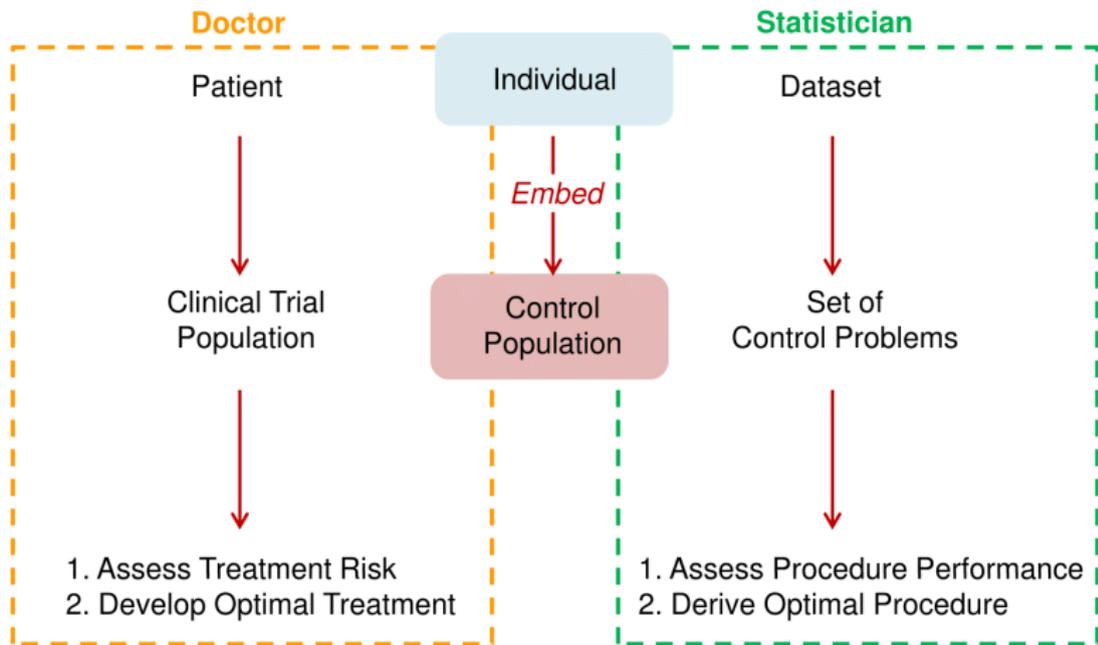
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# The Inevitable Statistical "Bootstrap": Creating Internal Replications

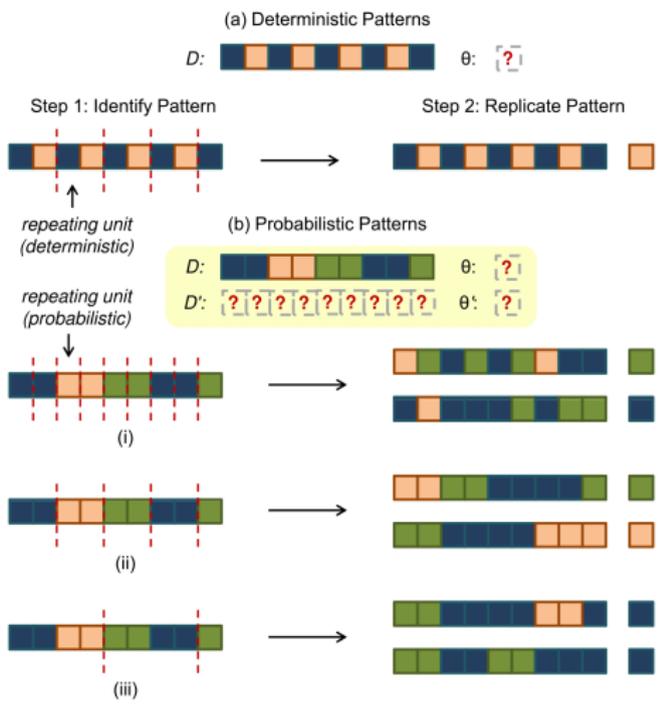
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# Relevant Controls/Replications are always needed

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Method Type	Error on Actual Problem $\Delta$	Average Error over Relevant Controls $\bar{\Delta}'$	References
<p><b>Point Estimate</b></p> <p><i>Goal:</i> Give our best guess, <math>\hat{\theta}</math>, for value of <math>\theta</math>.</p>	<p><math>L(\theta, \hat{\theta})</math></p> <p><i>Loss:</i> Specify how "far" <math>\theta</math> is from <math>\hat{\theta}</math> via loss function.</p>	<p><i>Risk:</i> The average loss of an estimator over control problems <math>(D', \theta')</math>.</p>	<p>Robinson 1979b Rukhin 1988, Lu and Berger 1989, Fourdrinier and Wells 2012</p>
<p><b>Set Estimate</b></p> <p><i>Goal:</i> Identify set, <math>C(D)</math>, of likely values for <math>\theta</math>.</p>	<p><math>I(\theta \notin C(D)) *</math></p> <p><i>Coverage:</i> Does our set contain the true value of <math>\theta</math>?</p>	<p><i>Non-Coverage Probability:</i> Proportion of times a set estimate, e.g. interval estimate, fails to contain the true value of <math>\theta'</math>.</p>	<p>Casella 1992, Goutis and Casella 1995, Robinson 1979a, Berger 1988</p>
<p><b>Hypothesis Test</b></p> <p><i>Goal:</i> Should we reject a null hypothesis, <math>H_0</math>, based on evidence from data?</p>	<p><math>I(\hat{T} \neq T)</math></p> <p><i>Type I or II Error:</i> Do we falsely reject or falsely accept <math>H_0</math>?</p>	<p><i>Error Probability:</i> The test's rates of false rejection and false acceptance when applied to control problems.</p>	<p>Hwang et al. 1992, Berger et al. 1994, Berger 2003</p>

\*  $I(\text{statement})$  denotes the indicator function: it equals 1 if the statement in parentheses is true and 0 otherwise.

# Multi-resolution Replications

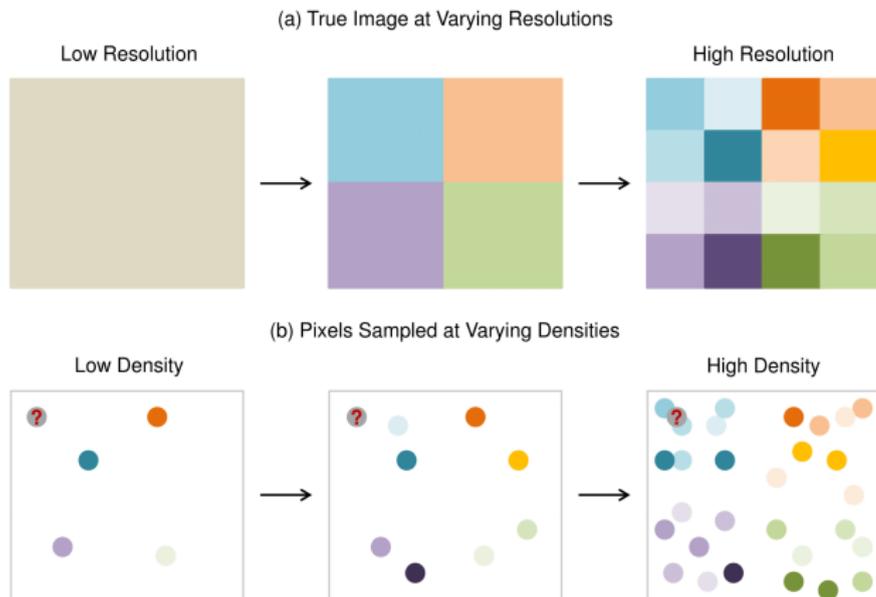
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# The Problem Gets Easier But My Intervals Get Longer ?!

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# When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Summary

## Precision as Function of Multiple Features (Basu 1964)

$(X_i, Y_i)$  bivariate standard normal with unknown correlation  $\theta$

# When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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Summary

## Precision as Function of Multiple Features (Basu 1964)

$(X_i, Y_i)$  bivariate standard normal with unknown correlation  $\theta$

- Fact 1:  $X_i, Y_i$  marginally ancillary, not jointly ancillary.

# When Ancillary Statistics Are Not Enough For Uncertainty Quantification

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$(X_i, Y_i)$  bivariate standard normal with unknown correlation  $\theta$

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**Option 1:** Evaluate uncertainty of  $\hat{\theta}$  (MLE) *unconditionally*.  
Construct pivot (using inverse CDF) and invert into CI.

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- Achieves **exact, unconditional** coverage.

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**Option 1:** Evaluate uncertainty of  $\hat{\theta}$  (MLE) *unconditionally*.  
Construct pivot (using inverse CDF) and invert into CI.

- Achieves **exact, unconditional** coverage.

**Option 2:** Evaluate uncertainty of  $\hat{\theta}$  *conditional* on  $\|X\|$ .

- But what about the effect of  $\|Y\|$  on precision?

# A Heterogeneous Population of Datasets

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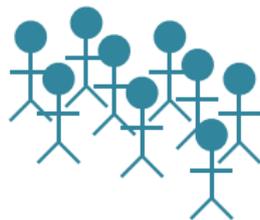
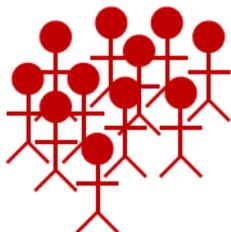
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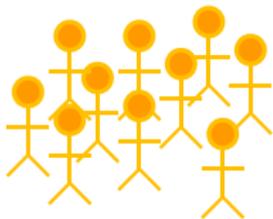
Low  $\|X\|$

High  $\|X\|$

Low  $\|Y\|$



High  $\|Y\|$



## A Regression Perspective

- As  $\|X\|$  increases, precision of  $\hat{\theta}$  increases.

# Here's Where Resolution Helps Us Reason...

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## A Regression Perspective

- As  $\|X\|$  increases, precision of  $\hat{\theta}$  increases.
- As  $\|Y\|$  increases, precision of  $\hat{\theta}$  increases.

# Here's Where Resolution Helps Us Reason...

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14/21

Xiao-Li Meng

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Basu Ex

Summary

## A Regression Perspective

- As  $\|X\|$  increases, precision of  $\hat{\theta}$  increases.
- As  $\|Y\|$  increases, precision of  $\hat{\theta}$  increases.
- The **first order** effects of  $\|X\|$  and  $\|Y\|$  on precision are robust to assumptions about  $\theta$ .

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- How to account for first order effects while ignoring second order effects and do so in a *principled* way?

# Fiducial's Pivotal Idea (Fraser 68, Hannig 09)

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## God's $U$ Always Exists

Represent data as  $X = g(\theta; U)$  where  $U \sim p(U)$  is known.

$$\text{Normal : } \quad X = \theta + U \quad U \sim N(0, 1)$$

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2. Convert inference for  $U$  into inference for  $\theta$  by inverting  $X = g(\theta; U)$  to obtain  $\theta = h(U; X)$ :

$$\text{E.g. : } \quad \theta = X - U \sim N(X, 1).$$

# Fiducial Inference for Bivariate Normal

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$(X_i, Y_i)$  bivariate normal with mean 0, var 1 and correlation  $\theta$ .

- Reduce to sufficient statistics:  $S_1 = \sum_i (X_i + Y_i)^2$  and  $S_2 = \sum_i (X_i - Y_i)^2$ .

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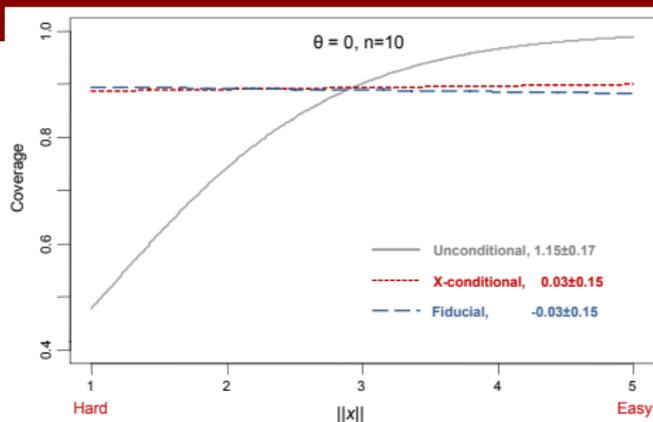
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- Inference for  $\theta$ : Given  $Q_1, Q_2$ , let  $\theta = \sqrt{\frac{S_1}{4Q_1}} - 1$ .

# Checking Coverage and Length Conditioning on

$\|x\|$



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$$\|x\|$$

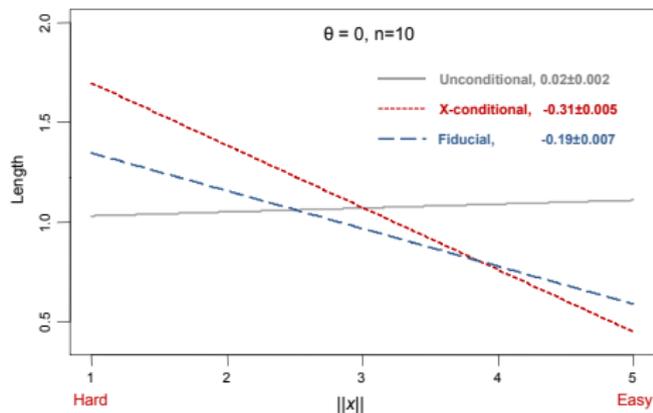
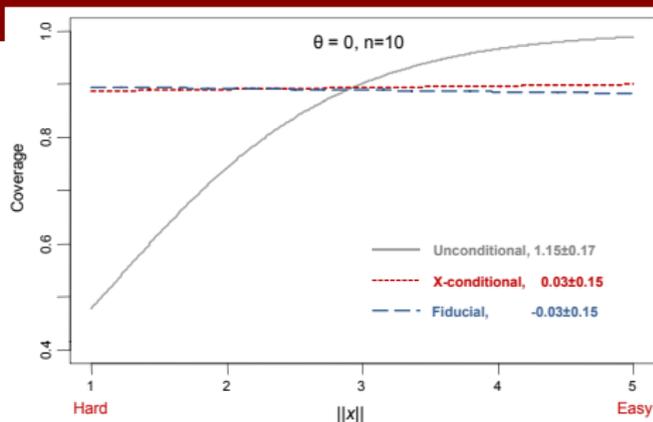
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# Checking Coverage and Length Conditioning on $\|y\|$

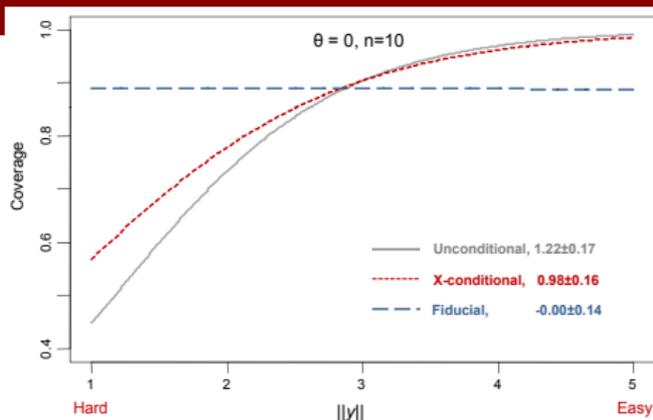
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## $\|y\|$

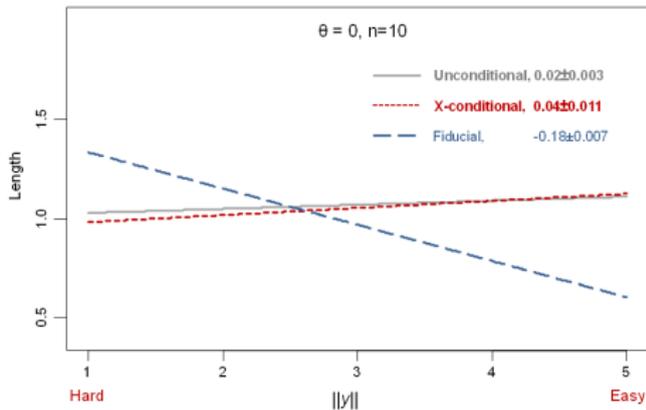
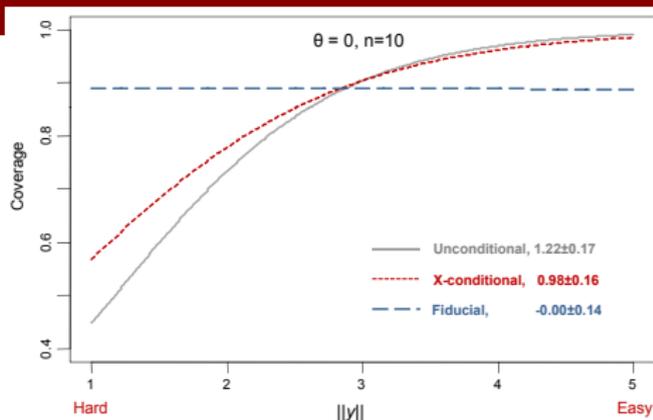
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# A Fundamental Principle of Statistical Inference: Bias-Variance or Relevant-Robust Trade-off

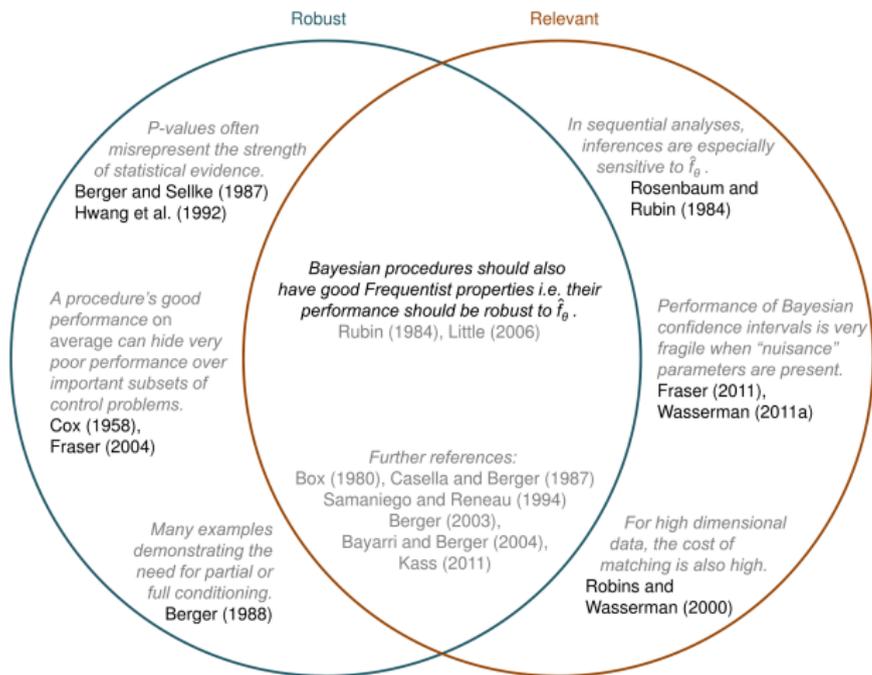
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# A Unified Picture of BFF (and Inference)?

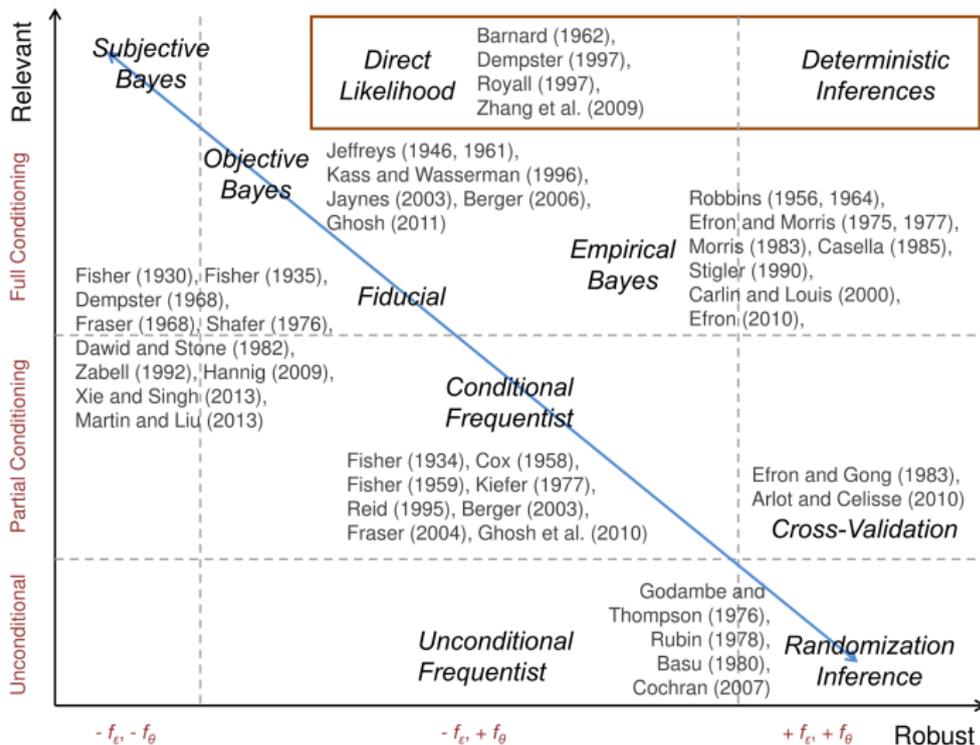
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# Let's be BFF, not merely FWB ...

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